

PSYCH 710

Higher-Order Within-Subjects ANOVA

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Factorial Within-Subject Design

- 2x3 within-subjects factorial design
- A & B are crossed, fixed factors
- subjects is a **random** factor
- typically, 1 observation per cell
 - not possible to measure within-cell variance
 - cannot distinguish contributions of error and subject x treatment interaction to between-cell variation
- data from each subject are correlated

	A ₁ B ₁	A ₁ B ₂	A ₁ B ₃	A ₂ B ₁	A ₂ B ₂	A ₂ B ₃
subject 1	n=1	n=1	n=1	n=1	n=1	n=1
subject 2	n=1	n=1	n=1	n=1	n=1	n=1
subject 3	n=1	n=1	n=1	n=1	n=1	n=1
subject 4	n=1	n=1	n=1	n=1	n=1	n=1

Full Model and F tests

$$Y_{ijk} = \mu + \alpha_j + \beta_k + \pi_i + (\alpha\beta)_{jk} + (\alpha\pi)_{ji} + (\beta\pi)_{ki} + (\alpha\beta\pi)_{jki} + \epsilon_{ijk}$$

- Statistical significance of all parameters evaluated by comparing $SS_{\text{residuals}}$ obtained with nested models
- error terms for F tests slightly more complicated:
 - main effect of A evaluated with A x Subjects term
 - main effect of B evaluated with B x Subjects term
 - A x B interaction evaluated with A x B x Subjects term

Effect	E(Mean Square)	F
S	$\sigma_\epsilon^2 + ab\sigma_\pi^2$	
A	$\sigma_\epsilon^2 + b\sigma_{\alpha\pi}^2 + nb\frac{\sum_j \alpha_j^2}{a-1}$	$\frac{MS_A}{MS_{A \times S}}$
A x S	$\sigma_\epsilon^2 + b\sigma_{\alpha\pi}^2$	
B	$\sigma_\epsilon^2 + a\sigma_{\beta\pi}^2 + na\frac{\sum_k \beta_k^2}{b-1}$	$\frac{MS_B}{MS_{B \times S}}$
B x S	$\sigma_\epsilon^2 + a\sigma_{\beta\pi}^2$	
A x B	$\sigma_\epsilon^2 + \sigma_{\alpha\beta\pi}^2 + n\frac{\sum_j \sum_k (\alpha\beta)_{jk}^2}{(a-1)(b-1)}$	$\frac{MS_{A \times B}}{MS_{A \times B \times S}}$
A x B x S	$\sigma_\epsilon^2 + \sigma_{\alpha\beta\pi}^2$	

R example

- Effects of noise (distractors) and stimulus orientation on letter discrimination
- 2 (noise) x 3 (orientation) within-subject factorial design

```
> load(url("http://pnb.mcmaster.ca/bennett/psy710/datasets/rtData-mw12-1.rda"))
> sapply(rt.wide,class)
      subj absent.a0 absent.a4 absent.a8 present.a0 present.a4 present.a8
"factor" "numeric"  "numeric"  "numeric"  "numeric"  "numeric"  "numeric"
> summary(rt.long)
      subj      noise      angle      rt
s1      : 6 absent :30 a0:20 Min.   : 284
s10     : 6 present:30 a4:20 1st Qu.: 459
s2      : 6                a8:20 Median : 571
s3      : 6                Mean   : 588
s4      : 6                3rd Qu.: 686
s5      : 6                Max.   :1000
(Other):24
```

R example

```
> options(width=80,digits=4)
> options(contrasts=c("contr.sum","contr.poly"))

> # following model is incorrect
> rt.aov.00 <- aov(rt ~ 1+Error(subj/(angle*noise)),
+               data=rt.long) # no fixed effects!
```

R example

```
> options(width=80,digits=4)
> options(contrasts=c("contr.sum","contr.poly"))

> # the next 2 models are equivalent:
> rt.aov.01 <- aov(rt ~ angle*noise+Error(subj/(angle*noise)),
+               data=rt.long)

> # rt.aov.01b <- aov(rt ~ angle*noise,
+                   + Error(subj/(angle+noise+angle:noise)),
+                   data=rt.long)
```

R example

```
> rt.aov.01 <- aov(rt ~ angle*noise+Error(subj/(angle*noise)),data=rt.long)
> summary(rt.aov.01)
Error: subj
      Df Sum Sq Mean Sq F value Pr(>F)
Residuals 9 344423  38269
```

```
Error: subj:angle
      Df Sum Sq Mean Sq F value Pr(>F)
angle   2 225422 112711  7.79 0.0036 **
Residuals 18 260280  14460
```

assumes sphericity

```
Error: subj:noise
      Df Sum Sq Mean Sq F value Pr(>F)
noise   1 234625 234625 15.7 0.0033 **
Residuals 9 134551  14950
```

```
Error: subj:angle:noise
      Df Sum Sq Mean Sq F value Pr(>F)
angle:noise 2 187983  93992  7.51 0.0043 **
Residuals 18 225402  12522
```

assumes sphericity

R example (aov_car)

```
> library(afex)
> rt.aov.02 <-aov_car(rt ~ angle*noise+Error(subj/(angle*noise)),data=rt.long)
> summary(rt.aov.02)
Univariate Type III Repeated-Measures ANOVA Assuming Sphericity
      Sum Sq num Df Error SS den Df F value Pr(>F)
(Intercept) 20765813      1  344423      9 542.62 2.4e-09 ***
angle        225423      2  260280     18   7.79 0.0036 **
noise       234625      1  134551      9  15.69 0.0033 **
angle:noise  187983      2  225402     18   7.51 0.0043 **
```

R example (aov_car)

```
> rt.aov.02 <-aov_car(rt ~ angle*noise+Error(subj/(angle*noise)),data=rt.long)
> summary(rt.aov.02)
Univariate Type III Repeated-Measures ANOVA Assuming Sphericity
      Sum Sq num Df Error SS den Df F value Pr(>F)
(Intercept) 20765813      1  344423      9 542.62 2.4e-09 ***
angle        225423      2  260280     18   7.79 0.0036 **
noise       234625      1  134551      9  15.69 0.0033 **
angle:noise  187983      2  225402     18   7.51 0.0043 **
```

Mauchly Tests for Sphericity

	Test statistic	p-value
angle	0.690	0.227
angle:noise	0.639	0.167

Greenhouse-Geisser and Huynh-Feldt Corrections

	GG eps	Pr(>F[GG])
angle	0.764	0.0083 **
angle:noise	0.735	0.0103 *

	HF eps	Pr(>F[HF])
angle	0.8880	0.005369
angle:noise	0.8431	0.007177

Strength of Association & Effect Size

omega-squared

$$\hat{\omega}^2 = \frac{df_{effect}(MS_{effect} - MS_{effect \times S})}{SS_{effect} + SS_{effect \times S} + SS_S + MS_S}$$

Cohen's f

$$\hat{f} = \sqrt{\frac{\hat{\omega}^2_{effect}}{1 - \hat{\omega}^2_{effect}}}$$

Strength of Association & Effect Size

```
library(effectsize)
omega_squared(rt.aov.02)

## # Effect Size for ANOVA (Type III)
##
## Parameter | Omega2 (partial) | 95% CI
## -----|-----|-----
## angle | 0.23 | [0.00, 1.00]
## noise | 0.29 | [0.00, 1.00]
## angle:noise | 0.20 | [0.00, 1.00]
##
## - One-sided CIs: upper bound fixed at (1).
```

Strength of Association & Effect Size

```
eta_squared(rt.aov.02)

## # Effect Size for ANOVA (Type III)
##
## Parameter | Eta2 (partial) | 95% CI
## -----
## angle | 0.46 | [0.14, 1.00]
## noise | 0.64 | [0.23, 1.00]
## angle:noise | 0.45 | [0.13, 1.00]
##
## - One-sided CIs: upper bound fixed at (1).
```

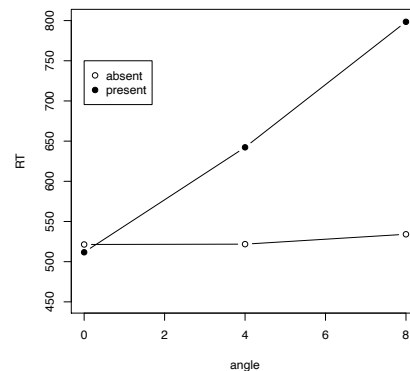
Strength of Association & Effect Size

```
cohens_f(rt.aov.02)

## # Effect Size for ANOVA (Type III)
##
## Parameter | Cohen's f (partial) | 95% CI
## -----
## angle | 0.93 | [0.41, Inf]
## noise | 1.32 | [0.55, Inf]
## angle:noise | 0.91 | [0.39, Inf]
##
## - One-sided CIs: upper bound fixed at (Inf).
```

simple main effects

- use simple main effects to decompose an interaction
- unlike between-subjects designs, it is better to use separate error terms for each simple main effect
- do not recalculate F with MS_{error} and df_{error} from main analysis
- each simple main effect is a 1-way within-subject ANOVA



simple main effects

```
aov.angle.absent <- aov_car(rt~angle+Error(subj/angle),data=subset(rt.long,noise=="absent"))
aov.angle.present <- aov_car(rt~angle+Error(subj/angle),data=subset(rt.long,noise=="present"))
# Huynh-Feldt correction; pes == partial eta squared:
anova(aov.angle.absent,correction="HF",es="pes")

## Anova Table (Type 3 tests)
##
## Response: rt
## num Df den Df MSE F pes Pr(>F)
## angle 2 18 10460 0.05 0.00551 0.95

anova(aov.angle.present,correction="HF",es="pes")

## Anova Table (Type 3 tests)
##
## Response: rt
## num Df den Df MSE F pes Pr(>F)
## angle 2 18 16522 12.5 0.581 4e-04 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

linear contrasts

- similar to procedures used with 1-way within-subjects ANOVA
- use contrast weights to create composite scores
 - converts multivariate analysis to univariate analysis
- use t test to evaluate null hypothesis

linear contrasts example

evaluate linear trend of RT across angle on entire data set (i.e., ignoring noise)

```
rt.mat <- as.matrix(rt.wide[,2:7])
rt.mat[1:2,]

##   absent.a0 absent.a4 absent.a8 present.a0 present.a4 present.a8
## 1      657      446      461      671      474      903
## 2      450      448      484      284      562      585

lin.C <- c(-1,0,1,-1,0,1)
rt.lin <- rt.mat %*% lin.C
```

linear contrasts example

evaluate linear trend of RT across angle on entire data set (i.e., ignoring noise)

```
lin.C <- c(-1,0,1,-1,0,1)
rt.lin <- rt.mat %*% lin.C
t.test(rt.lin)

##
##      One Sample t-test
##
## data:  rt.lin
## t = 3.5, df = 9, p-value = 0.007
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  104.9 494.1
## sample estimates:
## mean of x
##      299.5
```

linear contrasts example

evaluate linear trend of RT across angle separately at each level of noise

```
rt.mat[1,] # inspect column names

##   absent.a0 absent.a4 absent.a8 present.a0 present.a4 present.a8
##          657          446          461          671          474          903

# note the contrast weights in next 2 lines:
rt.absent.lin <- rt.mat %*% c(-1,0,1,0,0,0)
rt.present.lin <- rt.mat %*% c(0,0,0,-1,0,1)
```

linear contrasts example

evaluate linear trend of RT across angle separately at each level of noise

```
# note the contrast weights in next 2 lines:
rt.absent.lin <- rt.mat %*% c(-1,0,1,0,0,0)
rt.present.lin <- rt.mat %*% c(0,0,0,-1,0,1)
t.test(rt.absent.lin)

##
##      One Sample t-test
##
## data:  rt.absent.lin
## t = 0.32, df = 9, p-value = 0.8
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  -77.93 103.33
## sample estimates:
## mean of x
##      12.7
```

linear contrasts example

evaluate linear trend of RT across angle separately at each level of noise

```
# note the contrast weights in next 2 lines:
rt.absent.lin <- rt.mat %*% c(-1,0,1,0,0,0)
rt.present.lin <- rt.mat %*% c(0,0,0,-1,0,1)

t.test(rt.present.lin)

##
##      One Sample t-test
##
## data:  rt.present.lin
## t = 4.8, df = 9, p-value = 0.001
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  151.7 421.9
## sample estimates:
## mean of x
##      286.8
```

linear contrasts example

does linear trend of RT across angle differ across noise levels?

weights = $\Psi(-1,0,1,0,0,0) - \Psi(0,0,0,-1,0,1) = \Psi(-1,0,1,1,0,-1)$

```
myC <- c(-1,0,1,1,0,-1)          contrast weights (linear trend x noise interaction)
rt.lin.x.noise <- rt.mat %*% myC  composite scores
t.test(rt.lin.x.noise)           2-tailed t test

##
##      One Sample t-test
##
## data:  rt.lin.x.noise
## t = -5.1, df = 9, p-value = 7e-04      noise x linear trend interaction is significant
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  -396.8 -151.4
## sample estimates:
## mean of x
##      -274.1
```

linear contrasts example

using ANOVA to evaluate linear trend x noise interaction

```
N <- length(rt.absent.lin) # number of subjects
linTrend <- c(rt.pres.lin,rt.absent.lin)
nz <- as.factor(x=rep(c("present","absent"),each=N))
sid <- factor(x=rep(1:N,times=2),label="s")
linTrend.df <- data.frame(sid,nz,linTrend)
summary(linTrend.df)

##          sid          nz      linTrend
## s1         :2  absent :10  Min.      :-196
## s2         :2  present:10  1st Qu. :  19
## s3         :2                               Median : 104
## s4         :2                               Mean   : 150
## s5         :2                               3rd Qu.: 249
## s6         :2                               Max.   : 575
## (Other):8
```

create data frame of composite/trend scores for noise present & absent conditions

linear contrasts example

using ANOVA to evaluate linear trend x noise interaction

evaluate the effect of noise on composite scores with 1-way within-subjects ANOVA

```
options(contrasts=c("contr.sum","contr.poly"))
linTrend.aov.01 <- aov_car(linTrend-nz+Error(sid/nz),data=linTrend.df)
summary(linTrend.aov.01)
```

```
##
## Univariate Type III Repeated-Measures ANOVA Assuming Sphericity
##
##              Sum Sq num Df Error SS den Df F value Pr(>F)
## (Intercept) 448501     1  333077     9  12.1 0.00693 **
## nz          375654     1  132496     9  25.5 0.00069 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

what do these effects mean?

linear contrasts (emmeans)

evaluate linear trend of RT across angle on entire data set (i.e., ignoring noise)

```
library(emmeans)
rt.emm <- emmeans(rt.aov.02,specs="angle")
lin.C <- c(-1,0,1) # trend weights

# linear trend ignoring noise
contrast(rt.emm,method=list(linear=lin.C))

## contrast estimate SE df t.ratio p.value
## linear          150 43  9   3.481  0.0069
##
## Results are averaged over the levels of: noise
```

linear contrasts (emmeans)

evaluate linear trend of RT across angle separately at each level of noise

```
# linear trend separately for each noise
rt.emm.2 <- emmeans(rt.aov.02,specs="angle",by="noise")
contrast(rt.emm.2,method=list(linear=lin.C))
```

```
## noise = absent:
## contrast estimate SE df t.ratio p.value
## linear          12.7 40.1  9   0.317  0.7585
##
## noise = present:
## contrast estimate SE df t.ratio p.value
## linear          286.8 59.7  9   4.801  0.0010
```

linear contrasts (emmeans)

linear trend x noise interaction

```
# linear trend x noise interaction:
contrast(rt.emm.2,interaction=list(angle=list(lin.C),noise=list(c(-1,1))),by=NULL)

## angle_custom noise_custom estimate SE df t.ratio p.value
## c(-1, 0, 1) c(-1, 1)          274 54.3  9   5.051  0.0007
```

alternative methods (lmer)

```
# using lmer:
library(lmerTest)
rt.lmer.01 <- lmer(rt~angle*noise+(1|subj),data=rt.long)
anova(rt.lmer.01) # assumes sphericity

## Type III Analysis of Variance Table with Satterthwaite's method
##           Sum Sq Mean Sq NumDF DenDF F value Pr(>F)
## angle      225422  112711     2    45   8.18 0.00093 ***
## noise      234575  234575     1    45  17.02 0.00016 ***
## angle:noise 187983   93992     2    45   6.82 0.00259 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

assumes sphericity

alternative methods (lmer)

conservative F test

```
> # conservative F tests:
> N <- 10 # 10 subjects
> F.angle <- 8.18
> F.angle.x.noise <- 6.82
> # conservative F tests for angle and angle x noise:
> 1-pf(F.angle,1,N) # p value for angle
[1] 0.01695
> 1-pf(F.angle.x.noise,1,N) # p value for angle x noise
[1] 0.02597
```

alternative methods (lmer)

chi-square tests

```
> library(car)
> Anova(rt.lmer.01,type="III")
Analysis of Deviance Table (Type III Wald chisquare tests)

Response: rt
             Chisq Df Pr(>Chisq)
(Intercept) 542.6  1  < 2e-16 ***
angle        16.4  2  0.00028 ***
noise        17.0  1  3.7e-05 ***
angle:noise  13.6  2  0.00109 **
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

association strength (lmer)

fixed effects

```
library(effectsize)
omega_squared(rt.lmer.01)

## # Effect Size for ANOVA (Type III)
##
## Parameter | Omega2 (partial) | 95% CI
## -----|-----|-----
## angle | 0.23 | [0.06, 1.00]
## noise | 0.25 | [0.09, 1.00]
## angle:noise | 0.20 | [0.04, 1.00]
##
## - One-sided CIs: upper bound fixed at (1).
```


effect size (lmer)

fixed effects

```
cohens_f(rt.lmer.01)

## # Effect Size for ANOVA (Type III)
##
## Parameter | Cohen's f (partial) | 95% CI
## -----|-----|-----
## angle | 0.60 | [0.31, Inf]
## noise | 0.61 | [0.34, Inf]
## angle:noise | 0.55 | [0.25, Inf]
##
## - One-sided CIs: upper bound fixed at (Inf).
```

variance components (lmer)

```
> # random components
> ranova(rt.lmer.01) # anova-like table
```

ANOVA-like table for random-effects: Single term deletions

Model:

```
rt ~ angle + noise + (1 | subj) + angle:noise
      npar logLik AIC LRT Df Pr(>Chisq)
<none>      8 -350 715
(1 | subj)  7 -352 718 4.81 1 0.028 *
```

variance components (lmer)

```
> # variance components
> print(VarCorr(rt.lmer.01),comp=c("Variance","Std.Dev.))
Groups Name Variance Std.Dev.
subj (Intercept) 4081 63.9
Residual 13783 117.4

> # association strength
> library(performance)
> icc(rt.lmer.01,by_group = T)
# ICC by Group
Group | ICC
-----|-----
subj | 0.228
```

alternative methods (lme in nlme)

models differ in assumed correlation structure in residuals

```
> # using lme in nlme
> library(nlme)
> # assumes independence:
> rt.lme.01 <- lme(rt~angle*noise,random=~1|subj,
+ data=rt.long)
> # assumes sphericity:
> rt.lme.02 <- lme(rt~angle*noise,random=~1|subj,
+ data=rt.long,
+ correlation=corCompSymm(value=0.3,form=~1|subj))
> # does not assume sphericity:
> rt.lme.03 <- lme(rt~angle*noise,random=~1|subj,
+ data=rt.long,
+ correlation=corSymm(form=~1|subj))
```

alternative methods (lme in nlme)

models differ in assumed correlation structure in residuals

```
> # no significant difference in fit:
> anova(rt.lme.01,rt.lme.02,rt.lme.03)
      Model df   AIC    BIC logLik  Test L.Ratio p-value
rt.lme.01   1   8 715.5 731.4 -349.7
rt.lme.02   2   9 717.5 735.4 -349.7 1 vs 2   0.00  0.9999
rt.lme.03   3  23 728.3 774.0 -341.1 2 vs 3  17.23  0.2444
```

alternative methods (lme in nlme)

fixed effects ANOVA

```
# fixed effects for model 1
anova(rt.lme.01)

##                numDF  denDF  F-value  p-value
## (Intercept)         1     45    542.6  <.0001
## angle                2     45     8.2   0.0009
## noise                1     45    17.0   0.0002
## angle:noise         2     45     6.8   0.0026
```

alternative methods (lme in nlme)

fixed effects association strength

```
omega_squared(rt.lme.01)

## # Effect Size for ANOVA
##
## Parameter | Omega2 (partial) |          95% CI
## -----|-----|-----
## angle     |                   | [0.06, 1.00]
## noise     |                   | [0.09, 1.00]
## angle:noise |                   | [0.04, 1.00]
##
## - One-sided CIs: upper bound fixed at (1).
```

alternative methods (lme in nlme)

fixed effects effect size

```
cohens_f(rt.lme.01)

## # Effect Size for ANOVA
##
## Parameter | Cohen's f (partial) |          95% CI
## -----|-----|-----
## angle     |                   | [0.31,      Inf]
## noise     |                   | [0.34,      Inf]
## angle:noise |                   | [0.25,      Inf]
##
## - One-sided CIs: upper bound fixed at (Inf).
```

variance components (lme in nlme)

model parameters fit with assumption that residuals were independent

```
# variance components (independence)
print(VarCorr(rt.lme.01),comp=c("Variance","Std.Dev. "))
## subj = pdLogChol(1)
##          Variance StdDev
## (Intercept) 4081    63.89
## Residual    13783   117.40

(icc.subj <- 4081/(4081+13783) )

## [1] 0.2284
```

variance components (lme in nlme)

model parameters fit with compound symmetric variance-covariance matrix for residuals

```
# variance components (compound symmetry)
print(VarCorr(rt.lme.02),comp=c("Variance","Std.Dev. "))

## subj = pdLogChol(1)
##          Variance StdDev
## (Intercept)  1142    33.79
## Residual    16722   129.31

(icc.subj <- 1142/(1142+16722) )

## [1] 0.06393
```

split-plot (between-within, mixed) designs

- split-plot designs have between-subject & within-subject factors
- analyzed the same way as within-subjects design except we include between-subjects factors in multivariate linear model

split-plot designs

```
> library(afex)
> options(contrasts=c("contr.sum", "contr.poly"))
> rtAge.aov.02 <- aov_car(rt ~ group*angle + Error(subj/angle),data=rtVisual)
> summary(rtAge.aov.02)
```

Univariate Type III Repeated-Measures ANOVA Assuming Sphericity

	Sum Sq	num Df	Error SS	den Df	F value	Pr(>F)
(Intercept)	9610000	1	100222	10	958.87	2.9e-11 ***
group	16044	1	100222	10	1.60	0.234
angle	28850	2	114311	20	2.52	0.105
group:angle	50772	2	114311	20	4.44	0.025 *

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

split-plot designs

```

Sum Sq num Df Error SS den Df F value Pr(>F)
(Intercept) 9610000 1 100222 10 958.87 2.9e-11 ***
group 16044 1 100222 10 1.60 0.234
angle 28850 2 114311 20 2.52 0.105
group:angle 50772 2 114311 20 4.44 0.025 *

```

Mauchly Tests for Sphericity

```

Test statistic p-value
angle 0.354 0.00933
group:angle 0.354 0.00933

```

Greenhouse-Geisser & Huynh-Feldt Corrections

```

GG eps Pr(>F[GG])
angle 0.607 0.135
group:angle 0.607 0.051
HF eps Pr(>F[HF])
angle 0.6468 0.13155
group:angle 0.6468 0.04726

```

sphericity assumption applies to all effects that include a within-subjects factor (i.e., even between x within interactions)

split-plot designs

between-group portion is equivalent to 1-way ANOVA on average scores for each S

```

##
## Univariate Type III Repeated-Measures ANOVA Assuming Sphericity
##
##          Sum Sq num Df Error SS den Df F value Pr(>F)
## (Intercept) 9610000 1 100222 10 958.87 2.9e-11 ***
## group 16044 1 100222 10 1.60 0.234
## angle 28850 2 114311 20 2.52 0.105
## group:angle 50772 2 114311 20 4.44 0.025 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

rt.mat <- rtVisual.wide[,3:5] # extract rt measures
rtVisual.wide$rtAvg <- rowMeans(rt.mat) # calculate average
summary(aov(rtAvg~group,data=rtVisual.wide))

##          Df Sum Sq Mean Sq F value Pr(>F)
## group 1 5348 5348 1.6 0.23
## Residuals 10 33407 3341

```

split-plot designs

within-group parts of ANOVA use error term that is average of error terms in separate 1-way ANOVAs

```

##          Sum Sq num Df Error SS den Df F value Pr(>F)
## (Intercept) 9610000 1 100222 10 958.87 2.9e-11 ***
## group 16044 1 100222 10 1.60 0.234
## angle 28850 2 114311 20 2.52 0.105
## group:angle 50772 2 114311 20 4.44 0.025 *

```

```
summary(rt.young)[[2]]
```

```

##          Df Sum Sq Mean Sq F value Pr(>F)
## angle 2 72078 36039 9.71 0.0045 **
## Residuals 10 37122 3712

```

```
summary(rt.old)[[2]]
```

```

##          Df Sum Sq Mean Sq F value Pr(>F)
## angle 2 7544 3772 0.49 0.63
## Residuals 10 77189 7719

```

$$SS_{\text{error}} = 114311 = 37122 + 77189$$

$$MS_{\text{error}} = 5716 = (3712+7719)/2$$

linear contrasts

- on between-subject variable:
 - calculate mean score for each subject
 - apply contrast weights to mean scores as in a 1-way design
- on within-subject variable:
 - use contrast weights to convert measures to composite scores
 - use t-test or anova to determine if scores differ across groups (i.e., contrast x group interaction)

linear contrasts example

test of overall contrast ignoring group differences

```
> y.mat<-as.matrix( myData[,2:4] )
> lin.C <- c(-1,0,1)
> myData$lin.scores <- y.mat %>% lin.C
> myData
  group a1 a2 a3 lin.scores
1 young 50 47 51      1
2 young 41 57 43      2
3 young 42 63 40     -2
4 young 46 66 47      1
5 young 45 61 38     -7
6 young 45 57 53      8
7 old 48 39 38     -10
8 old 55 72 54     -1
9 old 51 44 51      0
10 old 53 65 53      0
11 old 68 58 62     -6
12 old 65 37 55     -10
```

mean contrast (linear trend) does not differ significantly from zero

```
> t.test(myData$lin.scores)

      One Sample t-test
data:  myData$lin.scores
t = -1.301, df = 11, p-value = 0.2199
alternative hypothesis: true mean is not equal to 0
95 percent confidence interval:
 -5.384  1.384
sample estimates:
mean of x
      -2
```

linear contrasts example

test of overall contrast ignoring group differences

```
> library(emmeans)
> # create aov_car object:
> rtAge.aov.10 <- aov_car(rt ~ group*angle + Error(subj/angle),
+                       data=rtVisual)
> linTrendWeights <- c(-1,0,1)
> # linear trend ignoring group
> angle.emm <- emmeans(rtAge.aov.10,specs="angle")
> contrast(angle.emm,method=list(linTrendWeights))
contrast estimate SE df t.ratio p.value
c(-1, 0, 1)    -20 14.1 10  -1.423  0.1851
```

Results are averaged over the levels of: group

linear contrasts example

evaluating group differences

```
> y.mat<-as.matrix( myData[,2:4] )
> lin.C <- c(-1,0,1)
> myData$lin.scores <- y.mat %>% lin.C
> myData
  group a1 a2 a3 lin.scores
1 young 50 47 51      1
2 young 41 57 43      2
3 young 42 63 40     -2
4 young 46 66 47      1
5 young 45 61 38     -7
6 young 45 57 53      8
7 old 48 39 38     -10
8 old 55 72 54     -1
9 old 51 44 51      0
10 old 53 65 53      0
11 old 68 58 62     -6
12 old 65 37 55     -10
```

contrast (linear trend) does not differ significantly between groups

```
> t.test(lin.scores~group,data=myData)

      Welch Two Sample t-test
data:  lin.scores by group
t = 1.779, df = 9.994, p-value = 0.1056
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
 -1.263  11.263
sample estimates:
mean in group young  mean in group old
              0.5             -4.5
```

linear contrasts example

evaluating group differences

```
> y.mat<-as.matrix( myData[,2:4] )
> lin.C <- c(-1,0,1)
> myData$lin.scores <- y.mat %>% lin.C
> myData
  group a1 a2 a3 lin.scores
1 young 50 47 51      1
2 young 41 57 43      2
3 young 42 63 40     -2
4 young 46 66 47      1
5 young 45 61 38     -7
6 young 45 57 53      8
7 old 48 39 38     -10
8 old 55 72 54     -1
9 old 51 44 51      0
10 old 53 65 53      0
11 old 68 58 62     -6
12 old 65 37 55     -10
```

contrast (linear trend) does not differ significantly between groups

```
> t.test(lin.scores~group,data=myData)

      Welch Two Sample t-test
data:  lin.scores by group
t = 1.779, df = 9.994, p-value = 0.1056
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
 -1.263  11.263
sample estimates:
mean in group young  mean in group old
              0.5             -4.5
```

linear contrasts example

evaluating group differences with emmeans

```
> # linear trend for each group
> angle.emm.2 <- emmeans(rtAge.aov.10,specs="angle",by="group")
> contrast(angle.emm.2,method=list(linTrendWeights))
group = young:
contrast estimate SE df t.ratio p.value
c(-1, 0, 1) 5 19.9 10 0.252 0.8065

group = old:
contrast estimate SE df t.ratio p.value
c(-1, 0, 1) -45 19.9 10 -2.264 0.0470

> # linear trend x group interaction
> contrast(angle.emm.2,interaction=c("poly","consec"),by=NULL)
angle_poly group_consec estimate SE df t.ratio p.value
linear old - young -50 28.1 10 -1.779 0.1056
quadratic old - young 307 143.2 10 2.142 0.0578
```